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# Statehood Experience, Legal Traditions and Climate Change Policies\*

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## Abstract

This paper investigates how the implementation of modern climate change policies is related to former colonies' length of state history and their legal heritage. We argue that countries with longer statehood experience around the time of colonization were better equipped to implement the legal philosophies transplanted by their colonial powers. Therefore, the implications of receiving British common law versus French civil law should be particularly important in countries with a greater accumulated history of statehood. Using a cross section of up to 78 former colonies, our results provide support for this hypothesis. In particular, our estimates demonstrate that common law countries have weaker modern climate change policies than civil law countries and the difference is inflated by a longer statehood experience, measured by the length of state history from 1-1800 AD. Legal origin has no effect in areas which, by the time of colonization, had no statehood experience.

*Key words:* Environmental policy; climate change; state antiquity; history; state capacity; legal origins; colonization.

*JEL classification:* Q58; K23; O44

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# 1 Introduction

Climate change has severe impacts on developing countries, affecting physical systems (rivers, lakes, floods, droughts, coastal erosion, sea level effects), biological systems (terrestrial and marine ecosystems), and human systems (health and livelihoods, food production) (IPCC, 2014). The literature explaining the pattern of policies addressing climate change and other environmental problems tends to emphasize relatively modern determinants, including the current levels of corruption, democracy, recent political instability, various political institutions, and trade policy.<sup>1</sup> However, history should not be ignored when we attempt to understand present-day policymaking. Is it possible that history going as far back as 1 AD still matters significantly for today’s environmental policy outcomes? In this paper, we study the joint effect of long-term statehood experience and legal heritage on climate change policy outcomes in former colonies. To the best of our knowledge, this is the first paper to merge these two strands of the literature. Moreover, it is the first to investigate the role of statehood history or state capacity for any type of environmental policymaking.

Long-term historical processes, including the development of agriculture, urbanization, the use of money as a medium of exchange, taxation, and experience with government administration all build stocks of human capital and experience through learning-by-doing (Burkett et al., 1999; Putterman, 2000; Acemoglu et al., 2015).<sup>2</sup> The level of accumulation of statehood experience influences a country’s ability to consolidate and centralize power, creates a capable bureaucracy which provides overall “state capacity” (Gennaioli and Rainer, 2007; Besley and Persson, 2009, 2011; Besley et al., 2013; Michalopoulos and Papaioannou, 2014). Besley and Persson (2011) present a theory predicting that state (fiscal) capacity and good institutions (legal capacity) are complements, which is supported by empirical evidence provided by Michalopoulos and Papaioannou (2014).

The form of legal system utilized has important implications for regulatory, judicial, financial, and economic outcomes; the literature has primarily focused on the implications of French civil law versus British common law for financial development and regulation (see, for example, La Porta et al., 1997, 1998, 2008; Botero et al., 2004). The resulting Legal Origins Theory (LOT) developed by La Porta et al. (2008) argues that British common law allocates more weight to private markets, the courts, and decentralized policymaking, while French civil law gives a strong role to the centralized state. Recent empirical work based on LOT finds that common law countries set weaker climate change policies than civil law countries (Fredriksson and Wollscheid, 2015).

So far, the literature on statehood experience has developed parallel to the literature on legal origins. We remedy this gap in the literature. We argue that the implications of receiving common law versus civil law should be particularly important in countries with a greater accumulated history of statehood at the time of colonization. In those countries, state (administrative) capacity was further developed, political power tended to be more centralized, the preexisting accountability of

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<sup>1</sup>See, for example, Congleton (1992), Damania et al. (2003), Fredriksson and Svensson (2003), Barbier et al. (2005), Farzin and Bond (2006), List and Sturm (2006), Hotte and Winer (2012), and Oliva (2015).

<sup>2</sup>The literature on the very long run determinants of economic development include, for example, Kremer (1993), Galor and Weil (2000), Comin et al. (2010) and Putterman and Weil (2010).

local leaders was greater, and the ability to receive and implement the transplanted legal framework from the colonizer was therefore enhanced. In contrast, areas without much statehood experience did not have a strong bureaucracy or state capacity, and the transfer of the colonizer’s legal approach was less complete. Our hypothesis is consequently that a longer history of statehood at the time of colonization serves to increase the differential impact of common and civil law legal origins on the stringency of modern climate change policies, as well as other environmental policies.

Using a cross section of up to 78 former colonies, our results lend support to this hypothesis. Our estimates suggest that the difference between common law and civil law countries in modern climate change policies is conditional on statehood experience, measured by the length of state history from 1-1800 AD. In particular, the divergence in modern climate change policy stringency due to legal origin (common law vs. civil law) is inflated by the length of statehood experience. Moreover, legal origin has no effect in countries which, at or around the time of colonization, had no statehood experience. We believe these are novel findings in the literature. Moreover, our findings apply more generally to several other measures of environmental policy outcomes.

These findings give a more fine-tuned picture of the role of legal origin compared to the previous literature on regulation in previous colonies. Our analysis helps improve our understanding of the long-term determinants of climate change policies and other environmental policies, as well as regulatory outcomes in general. Our result suggests that it is important to take the accumulated statehood experience into account when discussing the role of legal origin. The analysis may facilitate the prediction of which countries will tend to become leaders and laggards in the climate change policy formation process. An improved understanding of the institutional barriers created by legal and state history may also improve capacity building efforts, as well as the formulation of mechanisms and designs within international climate change treaties. Resources transfers may more easily be targeted towards their most productive uses.

The paper proceeds as follows. The next section reviews some of the related literature. Section 3 provides a discussion on the empirical specification, estimation issues and data. The empirical estimates are presented and analysed in Section 4. Several robustness checks are also performed, including the use of an alternative dependent variable. The last section summarizes and concludes.

## 2 Related Literature and Hypothesis Specification

The state is one of the most important forms of institutional development. The creation of nation states has led to a number of fundamental and far-reaching changes in human history. In recent years, statehood experience (sometimes denoted “state antiquity” in the literature) has gained considerable attention from the literature on long-run comparative economic development, helping to uncover the reasons for low income levels (Putterman, 2008; Putterman and Weil, 2010), bad institutions (Ang, 2013b), unequal distribution of income (Putterman and Weil, 2010), and financial underdevelopment (Ang, 2013a). The results of these studies generally suggest that a longer state history is associated with more favorable economic outcomes. A longer history of statehood often

implies a stronger state capacity in the form of legal and fiscal capabilities (Besley and Persson, 2009, 2011; Becerra et al., 2012). In particular, Michalopoulos and Papaioannou (2013) show that pre-colonial ethnic political centralization is strongly related to contemporary regional development in Africa. However, despite the above contributions, how statehood experience is related to environmental regulations has so far not been considered in the literature.

A large literature has investigated the implications of common law and civil law for modern economic and legal institutions, e.g., for the regulation of firm entry and labor; degree of judicial independence; formalism of judicial procedures; securities, company, and bankruptcy laws; and government ownership of banks (La Porta et al., 1997, 1998, 2008; Djankov et al., 2002, 2003; Botero et al., 2004). These institutions in turn have important implications for economic outcomes, e.g., unemployment and labor participation rates, corruption and the share of the unofficial economy, property rights, stock market development, and private credit. While legal systems were transplanted in a non-voluntary fashion to the colonies (David, 1985; McNeill and McNeill, 2003), legal philosophies are highly persistent even if domestic laws subsequently evolved in the recipient countries.

La Porta et al. (2008) present a Legal Origins Theory (LOT) which argues that British common law gives more weight to private markets and the courts, while civil law assigns an important role to the state.<sup>3</sup> Botero et al. (2004) suggest that countries with different legal origins essentially employ alternative institutional technologies for the social control of business activity (see also Glaeser and Shleifer, 2002). The legacies of legal origins exhibit a high degree of institutional persistence and path dependency where historical, economic, political, social and cultural features have a lasting impact (Greif, 1998; Hodgson, 1998; Acemoglu et al., 2001; North, 2006; Marchand, 2016).

Four implications of LOT are relevant for climate change policies. First, common law favors markets and judicial resolutions, not government regulation as emphasized by civil law.<sup>4</sup> Second, decentralized systems used in common countries are less prone to take transboundary pollution externalities into account compared to centralized systems emerging in civil law countries (see Sigman, 2014). Third, since common law respects private property relatively more, climate policies affecting capital owners are likely to be weaker in those countries. Fourth, while common law may be more adaptable to new situations as jurisprudence is more important as a source of law (Levine, 2005; Ponzetto and Fernandez, 2008), civil law country governments can more easily address emerging social problems due to a larger local bureaucracy (Mulligan and Shleifer, 2005). Overall, most of the above arguments support the hypothesis that common law should yield weaker climate change policies.

Fredriksson and Wollscheid (2015) report that former colonies with a common law legal heritage indeed set weaker climate change policies than civil law countries. Oto-Peralías and Romero-Ávila (2014) show that the effect of legal origins on several financial outcomes and business regulations

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<sup>3</sup>See Klerman and Mahoney (2007) and Roe (2007) for critical views of the Legal Origins Theory (LOT) of La Porta et al. (2008).

<sup>4</sup>Botero et al. (2004) show that British common law legal origin is associated with a substantial decrease in labor regulations.

is conditional on initial conditions (population density in year 1500 AD). In particular, the effect of common law is negatively related to pre-colonial population density, i.e. the outcome is closer to civil law countries when density is high. Berkowitz et al. (2003) argue that it does not matter significantly which form of law a country uses for it to be effective. According to them, what matters is whether the transplanted law was developed domestically, if it was received through colonization but adapted to local conditions, or whether the local population already had some familiarity with its legal principles. Under these circumstances, the demand for the law would be high and it would be used effectively to enhance economic development.

Gennaioli and Rainer (2007) discuss the importance of pre-colonial institutions in Africa, especially political centralization, for modernization programs during the colonial and the post-colonial periods. Policy coordination and implementation were facilitated by the existence of a pre-colonial hierarchy of chiefs in areas with centralized ethnic groups (see Schapera, 1970). Bargaining with senior traditional chiefs in these areas resulted in faster adoption of European policies than in areas with high political fragmentation. Preexisting accountability of local leaders in traditionally centralized systems was the mechanism which facilitated reforms. In fragmented areas, local chiefs pursued their own distorted policies, leading to lower rates of modernization and even tyranny (Tosh, 1978). Gennaioli and Rainer (2007) argue that institutions emerging in the African colonies were heavily influenced by the existing institutions found by the colonizers (see also Roberts and Mann, 1991).

Michalopoulos and Papaioannou (2013) document a strong association between African pre-colonial ethnic political centralization and modern regional development. Michalopoulos and Papaioannou (2013) argue that state capacity and legal capacity are complements in Africa, supporting the prediction of Besley and Persson (2011). In particular, they find that differences in national institutions (rule of law and corruption) yield differences in regional development among (Southern African) countries with high levels of state capacity. On the other hand, in (Western African) countries with low state capacity, no correlation is detected between national institutions and regional development.<sup>5</sup> These findings indicate that differences in the adoption of a colonizer's legal system across centralized and fragmented areas should similarly be determined by the level of statehood at the time of colonization. Path dependency and institutional persistence (see, e.g., Greif, 1998; North, 2006; Marchand, 2016) have preserved or inflated the resulting differences across legal origins over time.

Our hypothesis is thus that the transfer of legal origin should be more thorough and deep-rooted with a greater impact today in countries which at the time of colonization had a longer history of statehood experience. These colonies had a more developed legal and administrative bureaucracy, and hence were in a better position to facilitate the absorption of the legal philosophies. Therefore, the difference in the stringency of climate change policies between common and civil law countries should be greater in former colonies with a more extensive statehood experience. The divergence will increase due to a longer statehood history both because the anti-regulation stance of common

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<sup>5</sup>See Michalopoulos and Papaioannou (2013), Supplementary Appendix Table 16.

law will be more deeply embedded, and because the pro-regulation perspective of civil law becomes more deeply rooted in the society. Our hypothesis appears novel, as the previous literature has not investigated the regulatory effects of the absorption of legal origins conditional on state history, political centralization, or and other forms of state capacity.

### 3 Empirical Approach

#### 3.1 Regression model

The following model is regressed to investigate how climate change policy is related to statehood history and the legal tradition of a country:

$$CC_i^{policy} = \alpha + \beta State_i + \gamma_1 CommonLawLO_i + \gamma_2 MixedLawLO_i + \delta_1 State_i \times CommonLawLO_i + \delta_2 State_i \times MixedLawLO_i + cv'_i + \varepsilon_i \quad (1)$$

where  $CC^{policy}$  is an index of climate change policy,  $State$  is a measure of statehood experience covering the period 1–1800AD,  $CommonLawLO$  is a dummy variable for countries classified as having a British common law tradition,  $MixedLawLO$  is a dummy variable for countries classified as having a mixed (i.e., a combination of common and civil) law traditions,  $cv'$  is a set of control variables included in regressions to allow for the influence of some contemporary and geographic effects, and  $\varepsilon$  is an unobserved error term. “Other legal origins” is the excluded group, which includes 39 French and 3 German civil law countries.

Mixed legal tradition refers to those jurisdictions that have inherited both elements of common law and civil law. For instance, South Africa and Sri Lanka are coded as jurisdictions having a mixed legal tradition since they were initially colonized by the Netherlands and hence had inherited some form of civil law tradition. However, this legal system was partially replaced with common law when they were subsequently conquered by England (Klerman et al., 2011). Importantly, the mixed legal heritage arose due to exogenously determined events from the colonies’ perspectives.

As discussed in the previous section, our main variable of interest is the interaction term  $State \times CommonLawLO$ . The interaction term is expected to carry a negative sign due to the proposition that, relative to civil law countries, common law countries equipped with stronger state capacity impose less stringent climate change rules and regulations.

It is important to highlight that our estimations include only former colonies, which are not OECD members. This consideration is relevant since colonizers typically imposed a system of law upon their colonies. The legal structures imposed on former colonies under colonialism are therefore exogenously determined. By contrast, development of the legal systems in the origin countries, such as England, France, and Germany, was influenced by their own economic, cultural and political factors, and hence is endogenously determined.



Countries which adopted foreign legal systems voluntarily, such as Ethiopia, Japan, Thailand and Turkey, present similar endogeneity issues since they adopted a legal system, primarily French and German, which was deemed favorable for improving their economic conditions. Hence, a major advantage associated with the use of only former colonies in our analysis is that the exogenous nature of legal systems transplanted by colonialism enables us to interpret our estimates as reflecting the causal impact of legal tradition on climate policies (La Porta et al., 2008; Klerman et al., 2011)

### 3.2 Data

*Climate change policy index* ( $CC^{policy}$ ). We use the Climate Change Cooperation Index (C3-I) of Bernauer and Böhmelt (2013) as our measure of climate change policy strictness. The C3-I provides times-series cross-section data for up to 172 countries over the period 1996-2008. The index consists of a policy and an emissions component. The policy component, which is used as our main measure of climate change policy, is the aggregate of the following sub-components: (a) whether a country signed the United Nations Framework Convention on Climate Change (UNFCCC) and how soon this was done; (b) whether a country signed the Kyoto Protocol and how soon this was done; (c) whether a country submitted the latest national climate report and if this was done in time; and (d) how often a country made financial contributions to the UNFCCC on time over the period 1996-2008. These four sub-components are summed and an average value of the climate policy index ( $CC^{policy}$ ) is obtained for each country.

*State history* ( $State$ ). We use the latest version (version 3.1) of the state history data assembled by Putterman (2004), who provides state antiquity data covering 39 half centuries from 1 AD to 1950 AD for 151 countries. This index of state history gives a score from 0 to 50, reflecting: (1) the presence of a government above the tribal level (1 point if yes, 0 points if no); (2) whether this government is foreign or locally based (1 point if locally based, 0.5 points if foreign [i.e., the country is a colony], 0.75 if in between [a local government with substantial foreign oversight]; and (3) the proportion of the current territory covered by this government (1 point if over 50%, 0.75 points if between 25% and 50%, 0.5 points if between 10% and 25%, 0.3 points if less than 10%).

To illustrate, state history ( $State$ ) for the eighteen centuries to 1800 AD is calculated as follows:

$$State_i = \frac{\sum_{t=1}^{36} (1.05)^{1-t} \bullet S_{i,t}}{\sum_{t=1}^{36} (1.05)^{1-t} \bullet 50} \quad (2)$$

where  $S_{i,t}$  is the state presence for country  $i$  for the fifty-year period  $t$  (see Putterman and Weil, 2010). The scores on the three questions were multiplied by one another and by 50, so that for a given fifty year period, a country has a score of 50 if it was an autonomous nation, 0 if it had no government above the tribal level, 25 if the entire territory was ruled by a foreign country, and so on. Appendix 2 provides more details on the construction of State and several examples.

A 5 percent discount rate is applied to each of the half centuries so that less importance is attached to states formed in the more distant past. Our estimates are not sensitive to the use of alternative depreciation rates ranging from 0 to 20 percent. This approach of measuring state antiquity is broadly consistent with Putterman (2008), Putterman and Weil (2010), and more recently, Ang (2013a, b). The index is converted to a scale from 0 to 1 where higher values reflect the presence of a longer state history. Figure 1 shows the dispersion of State across the world for all available countries in the Putterman dataset.

[Figure 1]

Although data on state history are available up to 1950 AD, we measure it only up to 1800 AD – the period in which colonialism reached its peak. Doing so not only pre-empts the issue of reverse causality from climate policy to state capacity, but the period considered is also more in tandem with the timing when legal systems were imposed on colonies. In the robustness checks, state experiences accumulated up to 1500 AD and 1650 AD are also considered in order to check if the results are driven by the period chosen.

*Legal origins (LO)*. The legal tradition of company law or the commercial code for each country is classified into British common (*Common Law LO*) or Mixed (*Mixed Law LO*), with civil law as the excluded category, using binary variables. To do this, we follow the legal tradition classification of Klerman et al. (2011). The advantage of using this classification, as opposed to the more traditional classification approach of La Porta et al. (1998), is that it enables us to identify colonies which were influenced by both civil and common legal structures (Mixed), which may have some bearing on the results. Although Klerman et al. (2011) are legal scholars, who are expected to have detailed insights into countries legal histories, we also consider the more widely adopted legal system classification of La Porta et al. (2008) in the robustness analysis.

After assembling the various sources of data, 78 observations remain. The sample size is mainly constrained by the inclusion only of former colonies in the analysis. Table A2 provides a list of these countries, including legal origin and values of the *State* measure. Table 1 provides the summary statistics for the key variables used in the estimations.

[Table 1]

A simple comparison between common law and civil law countries with similar levels of State, such as Malaysia and Nigeria (common law;  $State = 0.585$  and  $0.544$ ;  $CC^{policy} = 36.923$  and  $30.082$ , respectively) vs. Armenia and Tajikistan (civil law;  $State = 0.561$  and  $0.524$ ;  $CC^{policy} = 49.279$  and  $53.267$ , respectively) indicates that common law countries tend to have weaker climate change policies. However, countries with no statehood experience in 1800 AD ( $State = 0.00$ ) such as Zambia (common law;  $CC^{policy} = 25.521$ ) and Gabon (civil law;  $CC^{policy} = 28.869$ ) exhibit only a small difference in  $CC^{policy}$ .

Figure 2 shows the overall association between climate change policy index and statehood experience. Figure 3 exhibits the various relationships between climate change policy index and

statehood experience across different types of legal tradition. Consistent with our prediction, these diagrams provide preliminary evidence that  $State$  and  $CC^{policy}$  are negatively correlated in countries that have the common law legal system. There is, however, no clear relationship between these variables for countries with “mixed” legal traditions.

[Figure 2]

However, it is necessary to check whether there is any difference in outcomes in colonies that did not have any statehood experience in 1800AD. In this connection, it is worth highlighting that the intercept values of the regression models are quite similar in all three sub-samples. This implies that countries which lacked any statehood experience in 1800 AD are likely to end up adopting more or less the same level of stringency in climate change policies. This empirical observation reinforces our proposition that how each type of legal tradition influences climate change policies depends crucially on statehood experience. To confirm this visual inspection, we perform an F-test on the equality of the intercept coefficients for the models that include common law countries only and civil law countries only (their respective intercept coefficients are 36.288 and 30.102). The test-statistic obtained from the nested regression model is 2.590 ( $p=0.113$ ), suggesting that the intercept values are not statistically different.

## 4 Results

### 4.1 Main findings

The estimation results of Eq. (1) are presented in Table 2. We consider several alternative specifications in which the main covariates are entered with different combinations in the regressions to ensure that the results are not driven by any particular model specification. In particular, column (1) considers only the effect of statehood whereas column (2) includes only the legal origin dummies. We do not find any evidence in support of the notion that the variation in climate change policy across countries can be solely accounted for by differences in statehood or legal traditions.

[Table 2]

Finally, the last column in Table 2 provides the full specification by adding common and mixed legal tradition and their interactions with state history. This complete specification will be used as our baseline model for robustness checks of the results. The results indicate that the interaction between statehood and the common law dummy is significantly correlated with climate change policy with the expected negative sign. This correlation is found to be statistically significant at the 1 percent level. Thus, the results suggest that a longer history of statehood experience at the time of colonization inflates the difference between civil law and common law countries’ stringency of climate change policy. The effects are economically meaningful.

The results in column (3) suggest that at the mean value of  $State$ , the presence of common law legal origin is associated with a change in  $CC^{policy}$  by  $(6.186 - 0.32 \times 28.069) = -2.796$ , equal to a

decline by 0.26 s.d.. At one s.d. above the mean of *State* (0.58), the marginal change in  $CC^{policy}$  equals -10.09, equivalent to 0.95 s.d. All other things being equal, this is equivalent to switching policy from the one in Gambia ( $CC^{policy} = 33.53$ ) to the one in Bangladesh ( $CC^{policy} = 23.44$ ). This may help explain why common law country Sudan, which had a relatively extensive state experience in 1800AD (*State* = 0.703; 2.20 s.d. above the mean), had  $CC^{policy} = 21.43$  (1.31 s.d. below the mean).

Statehood experience is found to be significantly correlated with climate change policy, with a positive sign consistent with our prediction. We also find that while the mixed legal system dummy variable is statistically insignificant, its' interaction with statehood is weakly significant in differentiating the variation in climate change policy observed across countries. On the whole, the results presented in Table 2 lend some initial (strong) support to our hypothesis that the stringency of climate change policy can be predicted by the strength of the interaction between statehood and common law legal heritage.

Figure 4 shows the partial regression line for the correlation between climate change policy and the interaction between state experience and common law system, while controlling for the influence of other variations stated in Eq. (1). Evidently, the partial regression line shows a strong negative correlation between the variables, thus reinforcing the findings in Table 2.

[Figure 4]

## 4.2 Controlling for other effects

To ensure that the above results are not confounded by the influence of some other effects and that the early advantage conferred by statehood is not proxying some forms of geographic characteristics, we include several control variables in the regressions and report the results in Table 3. The non-geographic variables include manufacturing output ratio, per capita real income, democratic capital, two indices of institutional quality and an index of ethnolinguistic fractionalization. Controlling for these effects is necessary since they may be correlated with statehood, legal origins and the outcome variable.<sup>6</sup> Geographic measures that have been argued to be potentially influential for economic outcomes including latitude, landlockedness, mean elevation, distance to the nearest coast or river, and precipitation. These are also controlled for in the regressions. Definitions and sources of these variables are given in Appendix 1.

[Table 3]

Columns (1) and (2) add the contemporary measures individually. Column (3) includes democratic capital stock. Column (4) controls for the effects of institutions (control of corruption and political stability), column (5) controls for ethnolinguistic fractionalization, and column (6) includes some additional effects of geography. The last column includes all control variables simultaneously.

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<sup>6</sup>One may suspect that statehood experience may be highly correlated with democratic capital and economic development. Their correlations are surprisingly low: -0.121 between *State* and GDP per capita and -0.138 between *State* and democratic capital.

It is evident that the coefficients of our core variable of interest, *State x Common Law LO*, vary only slightly from the baseline estimate reported in the last column of Table 2 when the effects of manufacturing output, democratic capital (discounted history of democracy), ethnolinguistic fragmentation and geography are allowed for in the regressions (columns (1), (3), (5) and (6)). In these models, the effect of *State x Common Law LO* remains precisely estimated and is at least significant at the 5 percent significance level.

In absolute terms, the coefficient of *State x Common Law LO* falls to 20.88 when per capita income is included as a control variable (column (2)). This result is not surprising given that prior studies have shown that state history and climate change policy are both significantly correlated with economic development. Similar findings are obtained when two measures of institutional quality are included as additional control variables in column (4). Nevertheless, the significant effect of *State x Common Law LO* prevails even after we control for the effect of income level or institutions.

In column (7), we add all control variables jointly into the specification. We find that *State x Common Law LO* continue to exert a strong explanatory power on climate change policy. This effect is found to be significant at the 5 percent level.

### 4.3 Alternative statehood periods

Next, we consider several alternative statehood periods to shed further light on the results. The estimations here include all the geographic controls used earlier, and are reported in Table 4. Columns (1) and (3) consider statehood only up to 1500 AD and 1650 AD, respectively. This is done to ensure that our results are not driven by states formed during the European colonial period since the 16th century. Interestingly, the results indicate that the absolute size of the coefficient on *State x Common Law LO* declines when the alternative dates are used, rather than the period 1-1800 AD as done previously in Tables 2 and 3. In particular, the coefficient sizes decline by 4 percent when statehood up to 1650 AD is considered, and by 8 percent when statehood up to 1500 AD is used (the estimates are benchmarked against column (6) of Table 3, which included the geographic controls). These findings imply that a longer statehood confers a stronger state capacity, which enables legal philosophies to develop deeper roots in society and the legal approach used has a greater policy impact. Importantly, the statistical significance of the *State x Common Law LO* interaction remains largely intact, suggesting that our main findings are unlikely to be driven by the formation of states due to the Western colonization.

[Table 4]

To ensure that the results are not driven by the exclusion of the remaining statehood period, we also add statehood 1501-1800 AD to the model in column (1) and statehood 1651-1800 AD to the specification in model (3). As shown in the results reported in columns (2) and (4), respectively, we obtain very similar findings.

In column (5), we measure statehood experience by considering only the presence of states for the period 1751-1800 AD. Doing so enables us to check whether it is the long-term statehood experience or the state presence at the peak of colonialism that matters. Interestingly, the estimates show that climate change policy strictness is unrelated to the state presence during this period. This finding therefore provides further credence to our proposition that the ability of a country to implement the legal philosophies transplanted by its colonial powers depends on the length of its statehood experience.

The results in column (6), which uses statehood from 1 to 1750 AD, confirm our previous finding that the length of statehood matters. In column (7), we show the results in column (5) are unchanged when state history up to 1750 AD is controlled for.

#### 4.4 Further Robustness Checks

Some further sensitivity checks are in order, and the results are reported in Table 5. First, we use the legal tradition classification of La Porta et al. (2008) to check if our results are robust to a coding method that is more widely used in the literature. The key difference between these two coding approaches relates to the classification of some jurisdictions as “mixed” by Klerman et al. (2011), which is not done by La Porta et al. (2008). Consequently, *Mixed Law LO* and *State x Mixed Law LO* are excluded from the regression for this sensitivity check. In our sample of 78 countries, only nine classifications are different. In Klerman et al. (2011), eight of these countries (i.e., Botswana, Cyprus, Guyana, Israel, Lesotho, South Africa, Sri Lanka and Zimbabwe) are classified as having a mixed system and one (i.e., Swaziland) as having adopted the French legal system, whereas La Porta et al. (2008) treat all of them as having British common law tradition. Hence, in principle, we should not expect much variation in the results since the common law dummy variables based on these two different coding procedures yield a correlation coefficient of 0.78. This conjecture is confirmed in the estimate reported in column (1).

[Table 5]

We include continent fixed effects in column (2) in order to ensure that the results are not being spuriously driven by unobserved time-invariant region-specific characteristics. Moreover, countries within the same continents tend to have similar colonization history and state performance, and these arbitrary correlations may bias our results. To address this concern, standard errors are clustered by continent to allow for these patterns within but not across continents. That is, the observations are assumed to be independent across continents but not within continents. The estimates in column (2), however, remain largely insensitive to this consideration.

Next, our results may be influenced by the presence of some influential observations such as Fiji, Korea and Uzbekistan. Since there is no compelling reason to exclude them from the estimations, we perform the robust regression analysis, which is a form of weighted regression technique that takes into account of the presence of extreme data points (see, e.g., Andersen, 2008). The results reported in column (3) indicate that our main findings prevail. We have also tried to separately exclude

the low statehood experience outlier (i.e., Fiji) and the two high statehood experience outliers (i.e., South Korea and Uzbekistan) from the estimations, but the coefficients of interest remain significant at the 5% level. When all three influential observations are removed, the coefficient of interest is still significant at the 10% level. The additional results are not reported to conserve space (available upon request).

Additionally, our results may potentially be driven by the failure to control for the initial conditions of colonization. In particular, civil law countries were previously colonized mainly by France, Portugal and Spain whereas common law countries by Great Britain. These countries developed different colonial models, based on differentials in endowment factors in the colonies, which had significant bearing on settlement decisions, institutional building (inclusive or exclusive), and consequently the legacies they left in the post-colonialism era (see also Sokoloff and Engerman, 2000; Acemoglu et al., 2001).

For instance, the Spanish colonizers mostly settled in the most politically and economically advanced regions (i.e., those with older states), whereas the British colonizers did largely the opposite. Consequently, the civil law variable may be a proxy for endowment factors in pre-colonial areas such as the statehood experience. In this connection, we do not find this conjecture supported by our data, given that statehood experience and legal origins are only weakly related (the correlation is only 0.15 between *State* and French civil law legal origin dummy).

Nevertheless, we cannot rule out the possibility that significantly different mean values of *State* exist between civil law and common law countries, thus reflecting the fact that endowments shape colonial institutional building. That is, if countries with a civil law tradition indeed have longer statehood experience than those with a common law origin due to the pursuit of different colonization strategy, then what we find would not reflect the effect of legal origins conditional upon the level of statehood experience, but rather the fact that statehood experience is longer in civil law countries due to different colonization strategy. We address this concern in the following ways.

First, in order to compare the mean values of *State* between civil law and common law countries, we perform the two-group mean-comparison test. This is essentially a t-test designed to compare means of the same variable between two groups. When the independent variable is chosen to be common law so that the two groups are common law ( $N = 24$ ) and non-common law ( $N = 54$ ) group, the t-statistic is 1.72 and the corresponding p-value is 0.09. When a comparison is made between French civil law countries versus others, the p-value is 0.189. On these grounds, we conclude that the difference of means in *State* between civil law and common law countries is statistically not different from 0.

Second, we directly control for endowment effects using the settler mortality measure of Acemoglu et al. (2001). Doing so, however, does not change the results in any significant manner (see column (4)). Note that the results are almost unchanged when the improved measure of settler mortality rate provided by Albouy (2012) is used (results unreported to conserve space; available upon request). Third, in order to account for the possibility that legal origins are shaped by colonial strategy, which may confound our results, we include four colonial origin dummies (British, French,

Portuguese and Spanish with other European as the excluded group). Our main results prevail (see column (5)).

We also provide some additional estimation by considering the individual components of the aggregate statehood experience index. Does one component of the index drive our results, or do all contribute? The overall index considered so far is based on three sub-components, which capture the presence of a government above the tribal level ( $ST^{PRESENCE}$ ), whether this government is foreign or locally based ( $ST^{AUTONOMY}$ ), and the proportion of the current territory covered by this government ( $ST^{COVERAGE}$ ). The presence of a state above the tribal level appears highly relevant as a measure of the administrative capacity to adopt the colonizer’s legal philosophy, for example. Appendix 2 provides the details of the individual components.

The results provided in the last three columns show that coefficients of the interaction between *State* and common law dummy are significant in all cases, suggesting that all dimensions of statehood experience matter. The interaction between  $ST^{COVERAGE}$  (which captures the extent to which the territory of the modern country under consideration was ruled by a particular government) and common law dummy is found to have the most significant correlation.<sup>7</sup>

#### 4.5 Endogeneity Issues

As is common in all empirical studies, it would be necessary to correct for the endogeneity of statehood experience due to measurement errors or omitted variable bias. In this case, it is expected that the interaction between statehood experience and the legal origin dummies are also endogenous regressors. In order to isolate the exogenous variation in statehood experience, we use the timing of agricultural transition as the instrument. The data are obtained from Putterman (2006).

This identification strategy is based on the well-known proposition of Diamond (1997) that agricultural transition significantly increased food supply, which enabled settled agricultural villages with small-scale political entities governed by supra-tribal authorities to compound into larger polities, leading to the emergence of fully-fledged states. Dramatic improvements in agricultural productivity also enabled polities to enhance their fiscal capacity through raising more tax revenues. Hence, the onset of the institutionalization of power relations triggered by the Neolithic transition was a key catalyst for state formation and the development of state capacity. This hypothesis suggests that an earlier transition to agriculture is expected to have a positive influence on the length of statehood experience.

The instrumental variable results and the first-stage regressions are presented in Table 6. It is evident that, in line with its OLS counterpart, the interaction term between statehood experience and common law legal origin is found to have a significant effect (columns (4)). In the first-stage regressions, the timing of agricultural transition is found to be a significant determinant of statehood experience, consistent with our prediction (column (1)). Similarly, the interaction of agricultural transition and legal origins significantly explains the interaction of statehood experience and legal origins (columns (2) and (3)).

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<sup>7</sup>Territorial integrity appears important for building bureaucratic capacity, possibly due to the associated stability.



[Table 6]

The first-stage F-statistics provide supportive evidence that agricultural transition is a suitable instrument. Given that multiple endogenous regressors are used, we also adopt the procedure of Shea (1997) for testing instrument relevance. This approach tests for the strength of the relationship between every endogenous variable and the excluded instruments, after partialing out the included instruments and other endogenous variables. A small value of the Shea partial R-squared would indicate that the instruments lack sufficient relevance to explain all the endogenous regressors. The Shea’s partial R-squared statistics range from 0.114 to 0.455, providing evidence that the instruments are quite strongly correlated with the endogenous variables, and hence are sufficiently relevant to explain the endogenous regressors.

Additionally, we also use the Anderson-Rubin (1949) method for testing the significance of the endogenous regressors in the structural equation. This approach is robust to weak instruments. The test rejects the null hypothesis that the coefficients of the excluded instruments are jointly equal to zero at the 5% level of significance, thus providing evidence that the endogenous regressors are relevant even in the presence of weak instruments.

To further increase the credibility of the results, we also perform the fractionally resampled Anderson-Rubin (*FAR*) test, using the procedure developed by Berkowitz et al. (2012), to test the joint significance of the endogenous regressors in our IV estimations. This weak exogeneity test is performed under the assumption that our instrument nearly, but not completely, satisfies the exclusion restriction. A rejection of the null implies that our instrument is endogenous. Clearly, the results indicate that the exclusion restriction assumption is not violated.

Although the IV estimates give consistent results, Bun and Harrison (2014) show that, under some general conditions, the OLS estimates of interaction terms are consistent and asymptotically normally distributed and that standard OLS inference is valid for their coefficients.<sup>8</sup> Given that OLS estimation and inference are reliable, the potential gains for using instruments are limited since the assumption of exclusion restriction cannot be fully satisfied. In light of this, the OLS estimator is our preferred approach.

#### 4.6 Alternative environmental policy measures

Our analysis has so far focused only on testing the interaction effect of statehood experience and legal origins on climate change policy. Our results indicate that common law countries with experienced statehood are less inclined to adopt stringent climate change policies. Can the same effect be observed on additional environmental policy outcomes, including those that address more local or regional pollution problems? We adopt the same regression framework and consider four other environmental protection variables.

First, in addition to the  $CC^{policy}$  index used above, Bernauer and Böhmelt (2013) also provide a more encompassing index of climate change policy outcomes,  $CCCI$ , which takes into consideration

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<sup>8</sup> These conditions include allowing for continuous and discrete interaction terms, correlation between endogenous and exogenous regressors, conditional heteroskedasticity and non-normality (Bun and Harrison, 2014).

the level and trend effects of emissions. We use this overall index as the dependent variable in column (1), Table 7. Note, however, that the  $CC^{policy}$  is our preferred measure since emissions are not necessarily policy-determined and are often affected by other hard-to-control-for socio-economic and political factors, and hence are more likely to subject our results to endogeneity bias. As a second measure of environmental policy stringency we use the (inverse) maximum lead content per gallon of gasoline in 1996 collected from Lovei, 1998) in column (2); Reyes (2015) discusses some health risks associated with lead exposure. Next, in column (3) we use the comprehensive environmental performance index of Hsu et al. (2014), which uses a broad spectrum of indicators to measure countries' protection of human health from environmental harm and the protection of ecosystems in year 2012. Finally, we also use a measure of the stringency of environmental regulations in the agricultural sector for year 1990 from Eliste and Fredriksson (2002) in the last column. This seeks to capture the overall regulatory framework (policies, legislation, and enforcement) addressing environmental problems associated with the agricultural sector.

[Table 7]

The results in Table 7 do not alter our main findings markedly given the fact that the coefficient of State x Common Law LO remains significant at conventional levels with a negative sign. These results suggest that our findings may apply also to more local or regional environmental problems.

## 5 Summary and Conclusions

A country's histories of, for example, agricultural development, urbanization, money, taxation, and government administration, all build stocks of human capital and experience. The cumulative experience with statehood influences a country's ability to consolidate power and create a capable bureaucracy, which may be summed up as "state capacity." We argue that the level of long term statehood experience should influence the propensity of former colonies to absorb and implement the legal frameworks transplanted by the colonial powers.

This paper tests the hypothesis that the influence of former colonies' legal heritage on modern climate change policies is conditional on their historical experience with statehood (state capacity). Our empirical work provides support for this hypothesis. The positive difference in the climate change policy stringency in former colonies with common law and civil law systems is enlarged by a longer history of statehood experience. Consistent with this finding, we also find that legal origin has no effect on countries which had no statehood experience around the time of colonization. We obtain similar results also for several alternative environmental policies. Our results suggest that long term historical processes should not be ignored when studying and designing modern policymaking, in particular those addressing climate change.

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**Table 1:** Descriptive statistics of key variables

Variable	Observations	Mean	Std. Dev.	Min.	Max.
<i>CC<sup>policy</sup></i>	78	35.35	10.64	14.49	62.74
<i>CC<sup>policy</sup> – common law LO</i>	24	33.65	9.13	18.75	56.77
<i>CC<sup>policy</sup> – mixed law LO</i>	12	34.23	10.12	14.49	52.16
<i>CC<sup>policy</sup> – other LO</i>	42	36.64	11.61	16.46	62.74
<i>State</i>	78	0.32	0.26	0.00	0.93
<i>State – common law LO</i>	24	0.24	0.27	0.00	0.81
<i>State – mixed law LO</i>	12	0.26	0.27	0.00	0.79
<i>State – other LO</i>	42	0.38	0.25	0.00	0.93
<i>Common Law LO</i>	78	0.31	0.46	0.00	1.00
<i>Mixed LO</i>	78	0.15	0.36	0.00	1.00
<i>Other LO</i>	78	0.54	0.50	0.00	1.00
<i>State x Common Law LO</i>	78	0.07	0.18	0.00	0.81
<i>State x Mixed Law LO</i>	78	0.04	0.14	0.00	0.79
<i>State x Other LO</i>	78	0.20	0.26	0.00	0.93
<i>Manufacturing (% GDP)</i>	76	13.34	7.69	2.89	38.67
<i>Income per capita (logged)</i>	76	7.14	1.46	4.82	10.62
<i>Democratic capital</i>	68	0.42	0.35	0.00	1.00
<i>Control of corruption</i>	78	40.98	26.39	1.46	97.57
<i>Political stability</i>	78	37.72	26.17	0.48	90.38
<i>Ethnolinguistic frac</i>	72	0.53	0.25	0.00	0.93
<i>Latitude</i>	78	0.22	0.15	0.00	0.67
<i>Landlocked</i>	75	0.31	0.46	0.00	1.00
<i>Elevation</i>	78	560.83	510.29	31.64	2672.87
<i>Distance to coast</i>	78	429.94	454.08	3.68	1875.13
<i>Precipitation</i>	78	95.10	65.63	2.91	259.95

*Notes:* The descriptive statistics provided in the table include up to 78 countries used in the baseline regressions. Sources and definition of data are described in the text and the data appendix.

**Table 2:** Main results

Dep. Var. = <i>CC<sup>policy</sup></i>	(1)	(2)	(3)
<i>State</i>	5.346 (1.027)		17.194** (2.188)
<i>Common Law LO</i>		-2.996 (-1.155)	6.186 (1.595)
<i>Mixed Law LO</i>		-2.409 (-0.714)	5.710 (1.355)
<i>State x Common Law LO</i>			-28.069*** (-2.914)
<i>State x Mixed Law LO</i>			-23.365* (-1.743)
R-squared	0.018	0.018	0.130
Observations	78	78	78

*Notes:* robust standard errors are used and *t*-statistics are reported in the parentheses. \*, \*\* and \*\*\* indicate significance at the 10%, 5% and 1% levels, respectively. The intercept estimate is not shown. The omitted legal origin group is “other legal origins”, which includes 39 French and 3 German civil law countries.



**Table 3:** Robustness analysis: including control variables

Dep. Var. = <i>CCpolicy</i>	(1)	(2)	(3)	(4)	(5)	(6)	(7)
<i>State</i>	12.99 (1.64)	14.03* (1.81)	13.09 (1.56)	15.71* (1.99)	13.66* (1.69)	15.94* (1.73)	6.30 (0.64)
<i>Common Law LO</i>	5.93* (1.68)	2.18 (0.54)	3.53 (0.66)	1.98 (0.45)	5.47 (1.35)	7.03 (1.61)	5.91 (1.05)
<i>Mixed Law LO</i>	2.23 (0.54)	3.97 (0.85)	3.58 (0.71)	3.26 (0.79)	3.86 (0.80)	5.54 (1.07)	0.97 (0.15)
<i>State x Common Law LO</i>	-27.13*** (-3.00)	-20.88** (-2.11)	-22.30** (-2.01)	-21.99** (-2.15)	-24.61** (-2.39)	-28.43** (-2.60)	-25.75** (-2.16)
<i>State x Mixed Law LO</i>	-17.43 (-1.30)	-25.56* (-1.75)	-19.56 (-1.24)	-26.22* (-1.76)	-20.68 (-1.41)	-21.42 (-1.52)	-20.08 (-1.11)
<i>Manufacturing (% GDP)</i>	0.54*** (3.03)						0.55*** (2.80)
<i>Income per capita (logged)</i>		2.50** (2.39)					2.07 (1.24)
<i>Democratic capital</i>			1.43 (0.26)				2.64 (0.55)
<i>Control of corruption</i>				3.79* (1.76)			4.29 (1.14)
<i>Political stability</i>				-0.55 (-0.28)			-0.71 (-0.27)
<i>Ethnolinguistic frac</i>					-6.98 (-1.30)		3.34 (0.48)
<i>Latitude</i>						0.03 (0.40)	0.09 (0.86)
<i>Landlocked</i>						-0.64 (-0.19)	-0.22 (-0.05)
<i>Elevation</i>						0.00 (0.99)	0.00 (0.75)
<i>Distance to coast</i>						0.00 (0.13)	0.01 (1.07)
<i>Precipitation</i>						-0.00 (-0.01)	-0.00 (-0.14)
R-squared	0.243	0.195	0.089	0.184	0.150	0.155	0.339
Observations	77	77	68	78	72	75	65

*Notes:* Robust standard errors are used and *t*-statistics are reported in the parentheses. \*, \*\* and \*\*\* indicate significance at the 10%, 5% and 1% levels, respectively. The intercept estimate is not shown. The omitted legal origin group is “other legal origins”, which includes 39 French and 3 German civil law countries. The following countries are excluded due to unavailability of data: Israel (column (1)), Zimbabwe (column (2)), Barbados, Grenada, New Zealand, Sao Tome and Principe, Canada, United States, South Korea, Australia, Israel and Seychelles (column (3)), Democratic Republic of the Congo, Kyrgyzstan, Sao Tome & Principe, Barbados, Seychelles and Grenada (column (5)), and Barbados, Grenada and Seychelles (column (6)). The last column includes all control variables and hence has only 65 observations.

**Table 4:** Alternative statehood periods

	(1)	(2)	(3)	(4)	(5)	(6)	(7)
Dep. Var. = <i>CC<sup>policy</sup></i>	State history up to 1500AD (State 1)	State history up to 1500AD (State 1) & 1501-1800AD (State 2)	State history up to 1650AD (State 1)	State history up to 1650AD (State 1) & 1651-1800AD (State 2)	State history 1751-1800AD (State 2)	State history up to 1750AD (State 1)	State history up to 1750AD (State 1) & 1751-1800AD (State 2)
<i>State 1</i>	12.960 (1.596)	9.893 (1.135)	14.299 (1.655)	11.603 (1.217)		15.503* (1.726)	15.409 (1.528)
<i>Common Law LO</i>	3.873 (1.091)	6.404 (1.097)	5.415 (1.395)	6.763 (1.045)	7.031 (0.994)	6.539 (1.565)	5.596 (0.808)
<i>Mixed Law LO</i>	2.601 (0.591)	9.099 (1.233)	3.929 (0.843)	8.757 (1.022)	2.917 (0.339)	5.046 (1.022)	3.321 (0.389)
<i>State 1 x Common Law LO</i>	- 26.194*** (-2.737)	-24.362* (-1.849)	- 27.307*** (-2.680)	-28.707* (-1.919)		-28.166** (-2.643)	-34.780** (-2.026)
<i>State 1 x Mixed Law LO</i>	-15.928 (-1.351)	-7.234 (-0.588)	-18.612 (-1.470)	-13.139 (-0.986)		-20.792 (-1.539)	-21.804 (-1.561)
<i>State 2</i>		6.660 (0.810)		5.888 (0.667)	0.186 (0.990)		0.025 (0.136)
<i>State 2 x Common Law LO</i>		-5.317 (-0.397)		-1.330 (-0.086)	-0.328 (-1.311)		0.124 (0.349)
<i>State 2 x Mixed Law LO</i>		-21.138 (-1.209)		-14.029 (-0.720)	-0.146 (-0.426)		0.098 (0.294)
R-squared	0.148	0.166	0.152	0.164	0.094	0.156	0.163
Observations	75	75	75	75	75	75	75
Geographic controls	YES	YES	YES	YES	YES	YES	YES

*Notes:* Robust standard errors are used and *t*-statistics are reported in the parentheses. \*, \*\* and \*\*\* indicate significance at the 10%, 5% and 1% levels, respectively. The intercept estimate is not shown. The omitted legal origin group is “other legal origins”, which include 39 French and 3 German civil law countries. The geographic controls included are latitude, landlockedness, mean elevation, distance to the nearest coast or river, and precipitation. The estimations exclude Barbados, Grenada, and Seychelles.

**Table 5:** Further robustness checks

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
	La Porta et al. data	Include continent dummies	Robust regression (for pres- ence of outliers)	Include settler mortality	Include colonial origin dummies	Using PRES- ENCE of <i>State</i>	Using AUTON- OMY of <i>State</i>	Using COVER- AGE of <i>State</i>
<i>State</i>	16.93* (1.79)	9.74** (3.83)	15.64* (1.84)	14.82 (1.46)	7.34 (1.05)	8.95** (3.29)	8.65** (3.19)	14.75** (4.39)
<i>Common Law LO</i>	6.54 (1.56)	2.48 (0.66)	6.31 (1.27)	6.07 (1.31)	-4.57 (-1.05)	6.19 (1.09)	5.70 (1.15)	7.27 (1.28)
<i>Mixed Law LO</i>		1.80 (0.43)	4.93 (0.76)	4.11 (0.68)	-5.93 (-1.27)	4.84 (1.25)	3.77 (0.99)	5.74 (1.66)
<i>State x Common Law LO</i>	-29.18*** (-2.72)	-22.47*** (-6.36)	-28.80** (-2.34)	-22.88* (-1.85)	-16.33* (-1.76)	-17.61* (-2.56)	-20.08** (-2.89)	-23.24** (-3.26)
<i>State x Mixed Law LO</i>		-16.56 (-1.75)	-20.31 (-1.30)	-22.26* (-1.95)	-10.27 (-0.83)	-12.30 (-1.47)	-13.99 (-1.69)	-16.13 (-1.68)
<i>Settler mortality</i>				-0.01** (-2.29)				
<i>Colonial origins (F-stat)</i>	—				8.03*** [ <i>p</i> = 0.000]			
R-squared	0.176	0.323	0.144	0.264	0.426	0.122	0.123	0.149
Observations	75	75	75	51	75	75	75	75
Geographic controls	YES	YES	YES	YES	YES	YES	YES	YES

*Notes:* Robust standard errors are used and *t*-statistics are reported in the parentheses. \*, \*\* and \*\*\* indicate significance at the 10%, 5% and 1% levels, respectively. The intercept estimate is not shown. The omitted legal origin group is “other legal origins”, which includes 39 French and 3 German civil law countries. The geographic controls included are latitude, landlockedness, mean elevation, distance to the nearest coast or river, and precipitation. The continent dummies are Asia, Africa, America and Europe (Oceania is the excluded group). Due to the unavailability of data, Barbados, Grenada, and Seychelles are excluded from all columns. Column (4) also excludes another 24 countries, including Armenia, Azerbaijan, Botswana, Cambodia, Cape Verde, Cyprus, Georgia, Israel, Jordan, Kazakhstan, Kyrgyzstan, Lesotho, Malawi, Mongolia, Mozambique, Philippines, Sao Tome and Principe, Swaziland, Syria, Tajikistan, Turkmenistan, Uzbekistan, Zambia and Zimbabwe.

**Table 6:** Addressing endogeneity issues

	(1)	(2)	(3)	(4)
	1 <sup>st</sup> -stage regression (Dep. Var. = State)	1 <sup>st</sup> -stage regression (Dep. Var. = State x Common Law LO)	1 <sup>st</sup> -stage regression (Dep. Var. = State x Mixed Law LO)	2 <sup>nd</sup> -stage regression (Dep. Var. = CC <sup>policy</sup> )
<i>Common Law LO</i>	-0.215** (-2.577)	-0.056 (-1.142)	-0.000 (-0.021)	13.874 (1.513)
<i>Mixed Law LO</i>	-0.016 (-0.156)	0.014 (0.733)	0.059 (0.759)	9.694 (1.233)
<i>Agr. transition</i>	0.045** (2.238)	0.002 (0.425)	0.001 (0.075)	
<i>Agr. transition x Common Law LO</i>	0.041** (2.314)	0.095*** (12.936)	0.000 (0.019)	
<i>Agr. transition x Mixed Law LO</i>	-0.021 (-1.137)	-0.002 (-0.786)	0.050*** (6.046)	
<i>State</i>				39.300 (1.380)
<i>State x Common Law LO</i>				-50.394** (-2.107)
<i>State x Mixed Law LO</i>				-30.112 (-1.261)
1 <sup>st</sup> -stage F-statistic for excluded instrument	19.16	55.88	12.98	-
Shea's partial R-squared	0.114	0.263	0.455	-
Anderson-Rubin Wald test	-	-	-	10.65 [p = 0.013]
Fractionally resampled Anderson-Rubin test				6.861 [p = 0.147]
R-squared	0.500	0.777	0.749	0.017
No. of obs.	73	73	73	73
Geographic controls	YES	YES	YES	YES

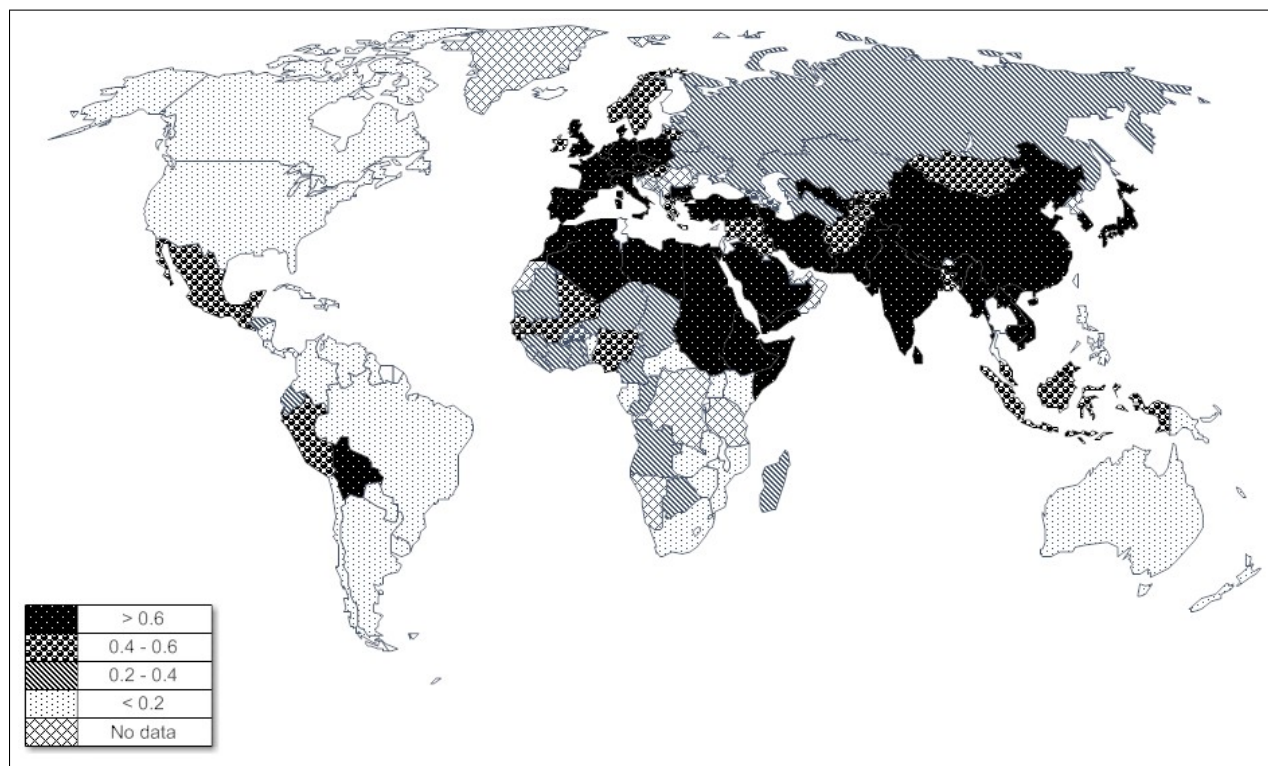
*Notes:* Robust standard errors are used and *t*-statistics are reported in the parentheses. \*, \*\* and \*\*\* indicate significance at the 10%, 5% and 1% levels, respectively. The intercept estimate is not shown. The omitted legal origin group is “other legal origins”, which includes 39 French and 3 German civil law countries. The geographic controls included are latitude, landlockedness, mean elevation, distance to the nearest coast or river, and precipitation. The estimations exclude Barbados, Grenada, Fiji, Sao Tome & Principe and Seychelles.

**Table 7:** Additional environmental policy outcomes

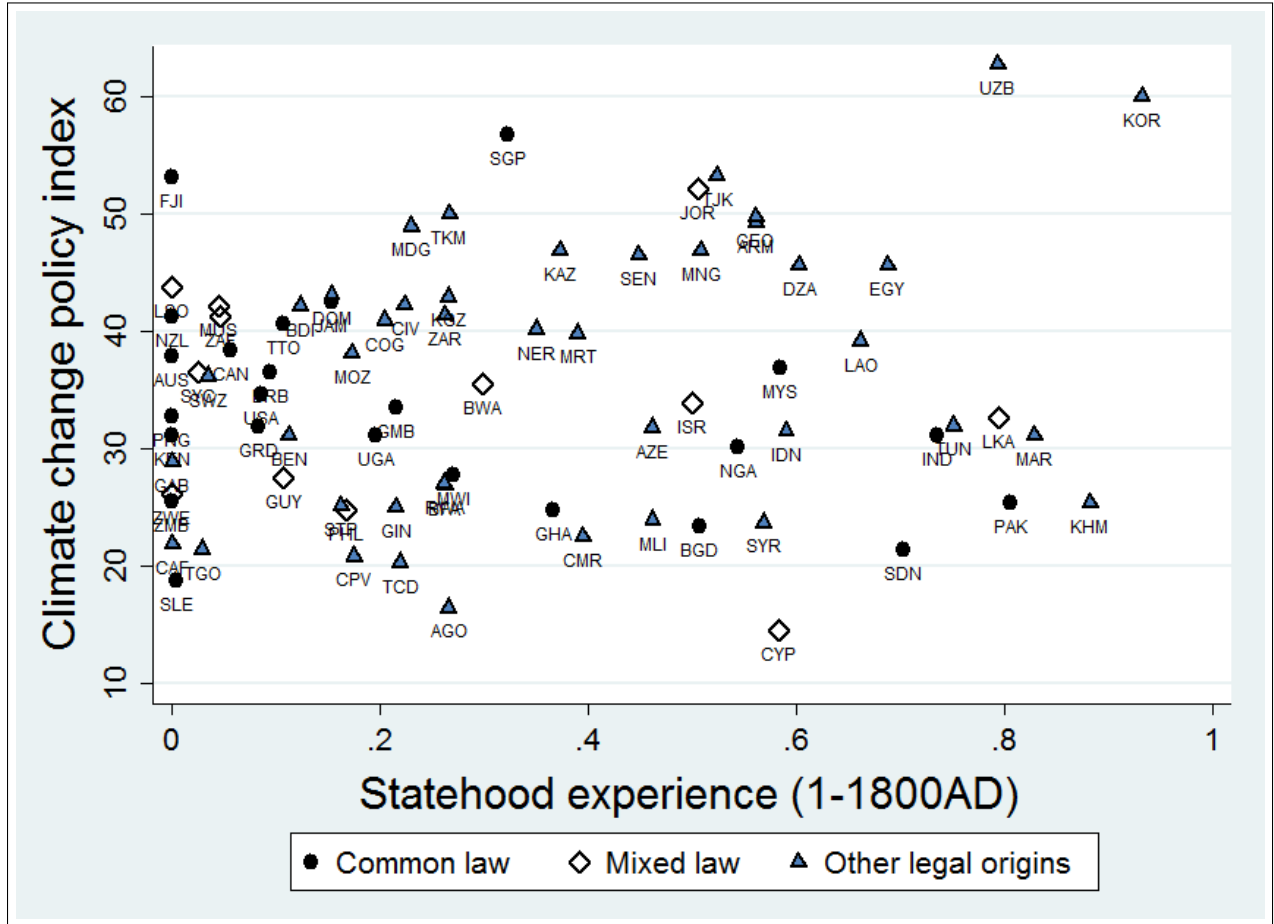
	(1)	(2)	(3)	(4)
Dep. Var. is:	<i>CCCI</i>	Max. lead content (inversed)	Environmental Performance Index	Environmental reg- ulation in agricul- ture
<i>State</i>	1.514 (1.344)	0.22 (0.81)	1.77 (0.41)	10.98 (0.21)
<i>Common Law LO</i>	0.537 (1.046)	0.39** (2.57)	8.46*** (2.78)	69.73*** (3.04)
<i>Mixed Law LO</i>	0.249 (0.483)	0.39*** (2.82)	0.23 (0.03)	60.62*** (3.86)
<i>State x Common Law LO</i>	-3.346*** (-2.665)	-0.56* (-1.71)	-21.67*** (-2.95)	-89.50* (-1.75)
<i>State x Mixed Law LO</i>	-1.174 (-0.840)	0.01 (0.02)	7.16 (0.57)	-151.59*** (-3.15)
R-squared	0.124	0.469	0.478	0.548
No. of obs.	74	43	52	26
Geographic controls	YES	YES	YES	YES

*Notes:* The alternative dependent variables are measured on different scales and hence sizes of the estimates are not directly comparable with earlier results or across table columns. Robust standard errors are used and *t*-statistics are reported in the parentheses. \*, \*\* and \*\*\* indicate significance at the 10%, 5% and 1% levels, respectively. The intercept estimate is not shown. The omitted legal origin group is “Other legal origins”, which includes 39 French and 3 German civil law countries. The geographic controls included are latitude, landlockedness, mean elevation, distance to the nearest coast or river, and precipitation.

**Figure 1:** Distribution of State (1 – 1800 AD) across the world

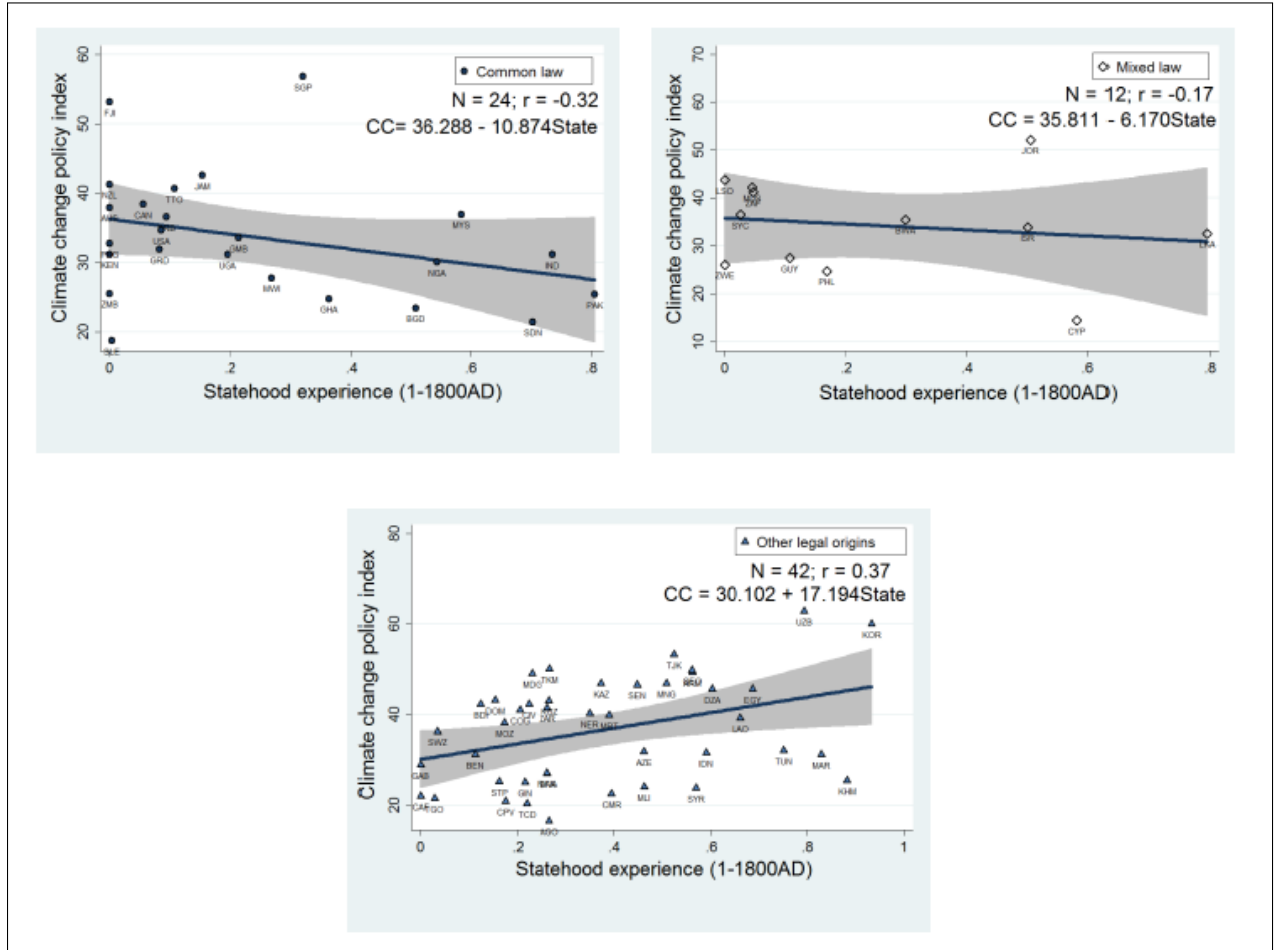


**Figure 2:** The correlation between Climate Change Policy Index ( $CC^{policy}$ ) and Statehood Experience ( $State$ )



*Notes:* The scatter plots include 78 countries used in the baseline estimations. “Other” legal origins include French (39 countries) and German civil law countries (3 countries).

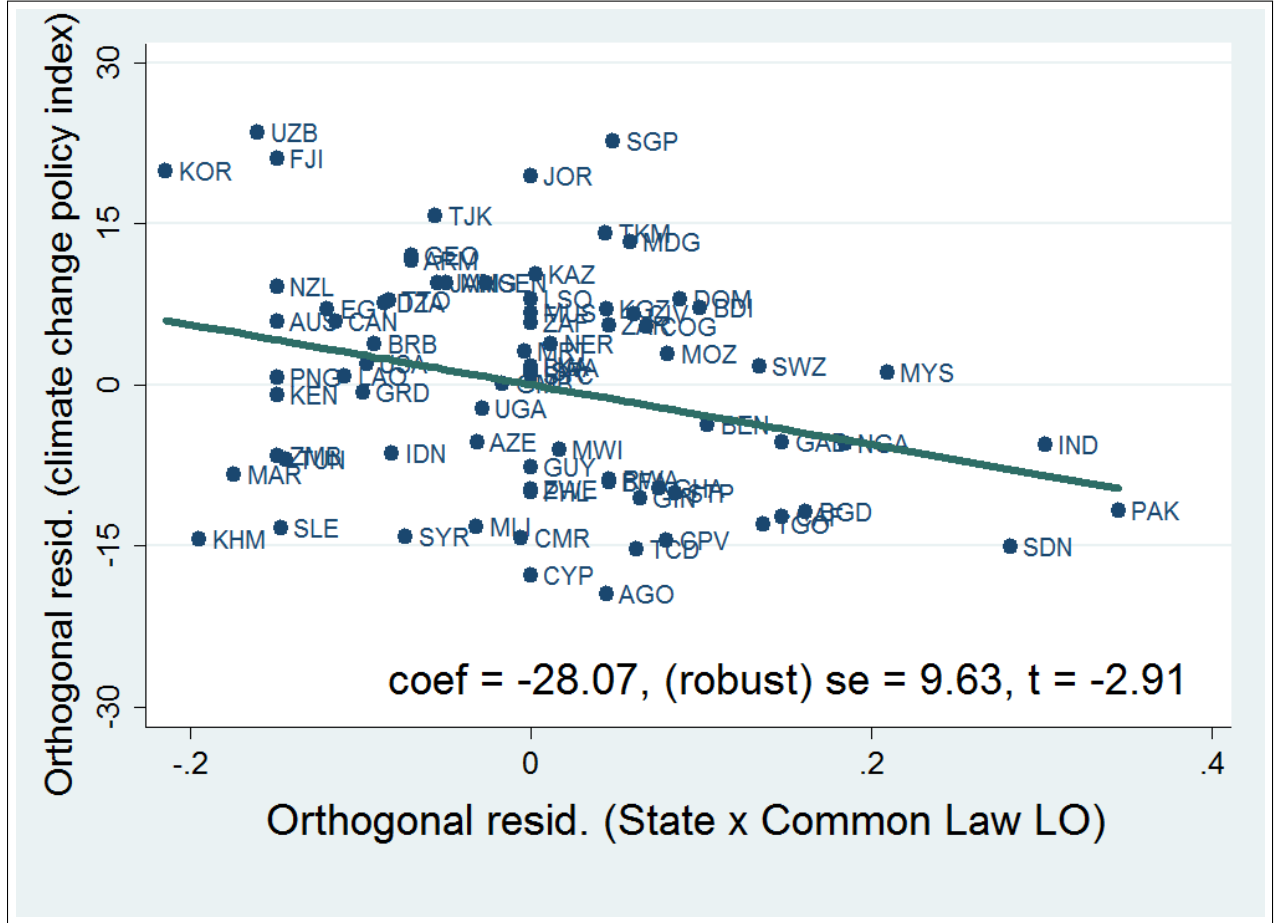
**Figure 3:** The relationships between Climate Change Policy Index and Statehood Experience across different legal traditions



*Notes:* “Other legal origins” includes the French and German civil law countries.  $N$  is the number of countries whereas  $r$  is the correlation coefficient between  $CC^{policy}$  and  $State$  in each type of legal tradition grouping. A 95 percent confidence interval band is used.



**Figure 4:** Partial effect of *State x Common Law LO*



*Notes:* The scatter plot in the above diagram illustrates the influence of State x Common Law LO on our climate change policy index while partialing out the effects of all other variables included in Eq. (1). This partial regression is based on the regression in column (3) in Table 2.

# Appendix 1: Data

**Table A1:** Definition of variables and data sources

Variable	Description	Source
<b>[A.] Main variables</b>		
<i>CC<sub>policy</sub></i>	A composite index capturing several key components of climate change policy, including the ratification of the United Nations Framework Convention on Climate Change (UNFCCC) and the Kyoto Protocol, the submission of national climate report, and financial contributions to the UNFCCC secretariat. The index also considers whether these activities were accomplished in a timely manner.	Bernauer and Böhmelt (2013)
<i>State</i>	An index of state history covering the period from 1 AD to 1800 AD, scaled to take values between 0 and 1. The latest version, v3.1, is used. In the robustness checks, two alternative periods, 1-1500 AD and 1-1650 AD, are used.	Putterman (2004)
<i>Common Law LO</i> <i>Civil Law LO</i> <i>Mixed Law LO</i>	A dummy variable that identifies the legal tradition of the company law or commercial code of each country as British common law, French civil law, German civil law or Mixed law. An alternative classification that excludes the “Mixed” category, by La Porta et al. (2008), is used in the robustness checks.	Klerman et al. (2011); La Porta et al. (2008)
<b>[B.] Control variables</b>		
<i>Manufacturing (% GDP)</i>	Manufacturing value added as a percentage of GDP (averaged over the period 1996-2008).	World Development Indicators (2015)
<i>Income per capita (logged)</i>	Logged GDP per capita converted to constant 2005 international dollars using PPP rates (averaged over the period 1996-2008).	World Development Indicators (2015)
<i>Democratic capital</i>	A measures of a country’s discounted stock of democracy from 1800 to 2010.	Fredriksson and Neumayer (2013)
<i>Control of corruption</i>	An index reflecting perceptions of the extent to which state power is used to obtain private gains (averaged over the period 1996-2008).	Worldwide Governance Indicators (2015)
<i>Political stability</i>	An index reflecting perceptions of the likelihood that the government will be destabilized or toppled (averaged over the period 1996-2008).	Worldwide Governance Indicators (2015)
<i>Latitude</i>	Value of the latitude of each country.	CIA (2015)
<i>Ethnolinguistic fractionalization</i>	An index capturing the probability that two randomly selected individuals from a country’s population would belong to different ethnic groups. The primary source used by Alesina et. al. (2003) is Encyclopedia Britannica (2000). The dates are generally from the early to mid-1990s (the range is 1979-2001, and the median year is 1994).	Alesina et al. (2003)
<i>Landlockedness</i>	A dummy variable that equals 1 if a country is fully enclosed by land and 0 otherwise.	CIA (2015)
<i>Distance to coast</i>	The mean distance of a country to the nearest coastline or sea-navigable river (in km)	Gallup et al. (2010)
<i>Elevation</i>	The mean elevation of a country above sea level (in km).	G-Econ (2008)
<i>Precipitation</i>	The average monthly precipitation of a country over the period 1961-1990 (in mm).	G-Econ (2008)
<i>Years since agricultural transition</i>	The number of years elapsed, in 2000 AD, since the transition to agriculture was estimated to occur (in thousand years).	Putterman (2006)
<i>Colonial origins</i>	A dummy variable that identifies the colonial origin of a country as Spanish, British, French, Portuguese or other European.	Nunn and Puga (2012)

<i>Settler mortality rate</i>	European settler mortality rates in the 19th century. Revised estimates of Albouy (2012) are also used as a check.	Acemoglu et al. (2001); Albouy (2012)
[C.] Other outcome variables		
<i>CCCI</i>	The aggregate climate change cooperation index of Bernauer and Böhmelt (2013). It combines <i>CC<sup>policy</sup></i> (our baseline measure of climate change policy) with data on the emission level and trend.	Bernauer and Böhmelt (2013)
Maximum lead content	The maximum lead content of lease gasoline in 1996 (gram per litre) permitted. The data are rescaled so that a higher value indicates more environmental protection.	Lovei (1998)
Environmental performance index	An environmental performance index for year 2012. Its computation covers the following nine policy dimensions: health impacts, air quality, water and sanitation, water resources, agriculture, forests, fisheries, biodiversity and habitat, and climate and energy.	Hsu et al. (2014)
Environmental regulation in agriculture	An index measuring the stringency of environmental regulation in agriculture in year 1990. Covers water pollution, air pollution, land use, and biodiversity.	Eliste and Fredriksson (2002)

**Table A2:** List of countries included in the estimations (by legal origin)

<b>Common Law LO</b>	<b>(avg = 0.243)</b>	Armenia	0.561	Senegal	0.448
Australia	0.000	Azerbaijan	0.462	Swaziland	0.035
Bangladesh	0.507	Benin	0.112	Syria	0.569
Barbados	0.094	Burkina Faso	0.262	Tajikistan	0.524
Canada	0.056	Burundi	0.123	Togo	0.029
Fiji	0.000	Cambodia	0.882	Tunisia	0.751
Gambia	0.216	Cameroon	0.395	Turkmenistan	0.267
Ghana	0.366	Cape Verde	0.175	Uzbekistan	0.794
Grenada	0.082	Central African Republic	0.000		
India	0.736	Chad	0.219	<b>Mixed Law LO</b>	<b>(avg = 0.256)</b>
Jamaica	0.153	Congo, Dem. Rep.	0.262	Botswana	0.299
Kenya	0.000	Congo, Rep.	0.205	Cyprus	0.583
Malawi	0.270	Cote d'Ivoire	0.223	Guyana	0.107
Malaysia	0.585	Dominican Republic	0.153	Israel	0.501
New Zealand	0.000	Egypt	0.688	Jordan	0.505
Nigeria	0.544	Gabon	0.000	Lesotho	0.000
Pakistan	0.806	Guinea	0.215	Mauritius	0.045
Papua New Guinea	0.000	Indonesia	0.590	Philippines	0.168
Sierra Leone	0.003	Kazakhstan	0.373	Seychelles	0.025
Singapore	0.322	Kyrgyzstan	0.265	South Africa	0.046
Sudan	0.703	Laos	0.662	Sri Lanka	0.795
Trinidad and Tobago	0.107	Madagascar	0.230	Zimbabwe	0.000
Uganda	0.196	Mali	0.462		
United States	0.086	Mauritania	0.390	<b>German Civil Law LO</b>	<b>(avg = 0.667)</b>
Zambia	0.000	Morocco	0.829	Georgia	0.561
		Mozambique	0.173	Korea, Rep.	0.933
<b>French Civil Law LO</b>	<b>(avg = 0.358)</b>	Niger	0.350	Mongolia	0.508
Algeria	0.603	Rwanda	0.262		
Angola	0.266	Sao Tome and Principe	0.162		

*Notes:* The estimations include up to 78 countries listed in the table above. Figures in the table indicate statehood experience accumulated over the period 1-1800 AD.

## Appendix 2: Construction of State History variable (*State*)

The state history index of Putterman (2004) includes 39 periods of 50 years spanning from 1 to 1950 AD. In our study, we restrict the period to 1-1800 AD (i.e., 36 periods). The index is made up of the following three components:

$ST^{PRESENCE}$ : Is there a government above the tribal level?  
[Yes = 1; No = 0]

$ST^{AUTONOMY}$ : Is this government foreign or locally based?  
[Local = 1; In between = 0.75; Foreign = 0.5]

$ST^{COVERAGE}$ : How much of the territory of the modern country was ruled by this government?  
[≥50% = 1; 25–50% = 0.75; 10–25% = 0.5; <10% = 0.3]

The extent of state presence ( $SA_t$ ) in any particular 50 years period ( $t$ ) is measured as the product of the scores on these components and 50. Consequently, a score of 0 indicates no presence of state, 25 reflects that a country has a supra-tribal authority but its entire territory is ruled by a foreign authority, and 50 indicates the presence of an autonomous nation, and so on.

$$SA_t = ST^{PRESENCE} \times ST^{AUTONOMY} \times ST^{COVERAGE} \times 50 \quad (A1)$$

$$0 \leq SA_t \leq 50, t = 1, 2, \dots, 36$$

The length of state history, or state antiquity (*State*), is measured as the cumulative presence of state by combining data over the entire 36 periods. A 5 percent discount rate is applied to allow for the fact that states formed in the more distant past have relatively less influence on today's economic conditions. To ease interpretation, the series is scaled into 0 and 1 using its maximum possible value. Accordingly, state history for a particular country over 18 centuries (1 – 1800 AD) is calculated as follows:

$$State = \frac{\sum_{t=1}^{36} (1.05)^{1-t} \cdot SA_t}{\sum_{t=1}^{36} (1.05)^{1-t} \cdot 50}, \quad 0 \leq State \leq 1 \quad (A2)$$

**Table A3:** Average values of *State* with sub-indices across continents

	<i>PRESENCE</i> <i>of State</i>	<i>AUTONOMY</i> <i>of State</i>	<i>COVERAGE</i> <i>of State</i>	<i>State</i>	No. of countries
Africa	0.419	0.371	0.332	0.282	46
Asia	0.866	0.688	0.786	0.619	37
America	0.364	0.233	0.325	0.193	27
Europe	0.829	0.609	0.770	0.551	37
Oceania	0.000	0.000	0.000	0.000	4
All countries	0.608	0.473	0.541	0.407	151

Notes: *PRESENCE* of *State* is measured as:  $(\sum_{t=1}^{36} (1.05)^{1-t} \cdot ST^{PRESENCE}_t \cdot 50) / (\sum_{t=1}^{36} (1.05)^{1-t} \cdot 50)$  and so on.  $ST^{PRESENCE}$  captures the presence of a government above the tribal level;  $ST^{AUTONOMY}$  reflects whether this government is foreign or locally based; and  $ST^{COVERAGE}$  measures the proportion of the current territory covered by this government. *State* is given as:  $(\sum_{t=1}^{36} (1.05)^{1-t} \cdot SA_t) / (\sum_{t=1}^{36} (1.05)^{1-t} \cdot 50)$ , where  $SA_t = ST^{PRESENCE} \times ST^{AUTONOMY} \times ST^{COVERAGE} \times 50$ . Data for all countries available are shown.

The distribution of *State* and its components for all available 151 countries is given in Table A3. In general, states in Asia and Europe are at least twice as old as those in Africa and America. Cambodia, China, Ethiopia, Japan, South Korea, Morocco, Myanmar, Nepal and Turkey were among the nations with the highest levels of state experience (average *State* = 0.883) by 1800 AD. Countries in Oceania,

including Australia, New Zealand, Fiji and Papua New Guinea had zero state values as of 1800 AD. African nations such as Central African Republic, Gabon, Kenya, Lesotho, Liberia, Seychelles, Sierra Leone, Togo, Zambia and Zimbabwe (average *State* = 0.006) are among those countries that have the youngest states, many of which also had no state experience at that time. It is interesting to observe that the emergence of states in several sub-Saharan African countries, such as the Central African Republic, Gabon, Kenya and Zimbabwe, was in fact the outcome of the 19th century colonization by the Western powers, either the British or the French, without which these nations may not have existed.

Tables A4 and A5 illustrate how the state history index is constructed for older and younger states, respectively, in each continent by giving a country example for each continent. Although our measure covers only 1-1800AD, for completeness we provide information for all time periods available (1-1950AD). The relatively experienced states considered are Egypt (Africa), China (Asia), Peru (America), France (Europe) and Australia (Oceania) whereas the relatively inexperienced states are Kenya (Africa), the Philippines (Asia), Uruguay (America), Latvia (Europe) and Papua New Guinea (Oceania).

**Table A4:** Examples of state history construction for experienced states in each continent

Region	AFRICA	ASIA	AMERICA	EUROPE	OCEANIA				
Country	Egypt	China	Peru	France	Australia				
1-50 AD	1, 0.5, 1 (under Roman's occupation)	1, 1, 1 (ruled by the Han dynasty)	0, 0, 0 (non-existence of states)	1, 0.5, 1 (Gaul was under the Roman's rule)	0, 0, 0 (non-existence of states)				
51-100 AD									
101-150 AD		1, 1, 0.75 (the Han empire was split into several warring states)		1, 1, 0.75 (controlled by multiple Germanic kingdoms)					
151-200 AD									
201-250 AD									
251-300 AD		1, 1, 0.5 (the Huari state existed in the southern part)							
301-350 AD									
351-400 AD		1, 1, 1 (unified under the Sui and, later, the T'ang dynasties)	1, 1, 1 (Huari grew to become an empire)	1, 1, 0.895 (rule was unified by Charlemagne in 771)					
401-450 AD									
451-500 AD		1, 1, 0.75 (central regime broke down, followed by some political chaos)	1, 1, 0.75 (the Huari empire collapsed and the area was ruled by a number of smaller states)	1, 1, 1 (under unified domestic rule)					
501-550 AD									
551-600 AD		1, 1, 1 (the T'ang power was restored)							
601-650 AD									
651-700 AD									
701-750 AD									
751-800 AD	1, 1, 1 (establishment of the Tulunid dynasty)	1, 1, 0.75 (centralized order collapsed again and multiple kingdoms emerged)		1, 1, 1 (under unified domestic rule)					
801-850 AD									
851-900 AD	1, 1, 1 (establishment of the Tulunid dynasty)	1, 1, 1 (power was united under the Sung dynasty)		1, 1, 0.75 (rule was divided among several Frankish kingdoms)					
901-950 AD	1, 0.5, 1 (became a foreign-based caliphal province)								
951-1000 AD	1, 0.77, 1 (the Fatimid Caliphate was established in 973)								
1001-1050 AD	1, 1, 1 (under the rule of the Egypt-based Fatimids)								
1051-1100 AD	1, 1, 0.75 (the Huari empire collapsed and the area was ruled by a number of smaller states)								
1101-1150 AD									
1151-1200 AD	1, 0.855, 1 (the Ayyubid dynasty became allegiant to the foreign-based Abbasid Caliphate in 1171)	1, 1, 1 (under the autonomy of the Mamluk dynasty)							
1201-1250 AD	1, 0.75, 1 (became quasi-independent under the Ayyubids)								
1251-1300 AD	1, 1, 1 (under the autonomy of the Mamluk dynasty)	1, 0.895, 1 (the Yuan dynasty was established in 1279)				1, 1, 1 (under centralized rule and area expanded to nearly its current size)			
1301-1350 AD		1, 0.75, 1 (under the quasi-local Mongol rule)							
1351-1400 AD		1, 0.91, 1		1, 0.75, 0.75					

		(the Yuan dynasty collapsed and replaced by the Ming dynasty in 1368)		(a large part of the area was held by the English during the Hundred Years War)	
1401-1450 AD		1, 1, 1 (the Chinese rule was integrated under the Ming and Ch'ing dynasties)			
1451-1500 AD			1, 1, 1 (the Inca empire united the entire area)	1, 1, 1 (the Hundred Years War ended and territory was regained)	
1501-1550 AD	1, 0.82, 1 (conquered by the Spanish in 1532)				
1551-1600 AD	1, 0.5, 1 (under the Spanish colonial rule)		1, 1, 0.75 (rule was divided by domestic religious wars)		
1601-1650 AD					
1651-1700 AD					
1701-1750 AD					
1751-1800 AD					
1801-1850 AD	1, 0.5, 1 (under the Ottomans' rule)		1, 0.75, 1 (allowed partial autonomy under the Ottomans and the French)		
1851-1900 AD	1, 0.66, 1 (under the British occupation in 1882)			1, 1, 1 (mostly under unified control)	
1901-1950 AD	1, 0.78, 1 (independent from Britain in 1922)	1, 1, 0.875 (The Ch'ing rule ceased in 1911 and was replaced the Republic of China, but the nation was politically fragmented)	1, 1, 1 (under independent rule)	1, 1, 1 (gained independence from Great Britain)	

Notes: The values in each entry reflect  $ST^{PRESENCE}$ ,  $ST^{AUTONOMY}$  and  $ST^{COVERAGE}$ . For instance,  $ST^{PRESENCE}$  is 1,  $ST^{AUTONOMY}$  is 0.5 and  $ST^{COVERAGE}$  is 1 for Egypt in every 50-year period from 1 to 850 AD. Source: Putterman (2004).

**Table A5:** Examples of state history construction for inexperienced states by continent

Region	AFRICA	ASIA	AMERICA	EUROPE	OCEANIA
Country	Kenya	Philippines	Uruguay	Moldova	Papua New Guinea
1-50 AD	0, 0, 0 (non-existence of states)	0, 0, 0 (non-existence of states)	0, 0, 0 (non-existence of states)	0, 0, 0 (non-existence of states)	0, 0, 0 (non-existence of states)
51-100 AD					
101-150 AD					
151-200 AD					
201-250 AD					
251-300 AD					
301-350 AD					
351-400 AD					
401-450 AD					
451-500 AD					
501-550 AD					
551-600 AD					
601-650 AD					
651-700 AD					
701-750 AD					
751-800 AD					
801-850 AD					
851-900 AD					
901-950 AD					
951-1000 AD					
1001-1050 AD					
1051-1100 AD					
1101-1150 AD					
1151-1200 AD					
1201-1250 AD					
1251-1300 AD					
1301-1350 AD					
1351-1400 AD					
1401-1450 AD					
1451-1500 AD					
1501-1550 AD					
1551-1600 AD					
1601-1650 AD					
1651-1700 AD					
1701-1750 AD					
1751-1800 AD					
1801-1850 AD					
1851-1900 AD					
1901-1950 AD	1, 0.5, 1 (became a colony of the British)	1, 0.75, 1 (local autonomy coexisted with control by the United States)	1, 1, 1 (under independent rule)	1, 0.5, 1 (largely occupied by Soviet Union)	1, 0.5, 0.75 (occupied by the British, Germans, and Dutch)

Notes: The values in each entry reflect  $ST^{PRESENCE}$ ,  $ST^{AUTONOMY}$  and  $ST^{COVERAGE}$ . Source: Putterman (2004).